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MARGINAL AND CONDITIONAL TREATMENT EFFECTS IN CLINICAL TRIALS

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on behalf of conditional vs marginal task force



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Objective:

We would like to bring the complex concept and methods about conditional and marginal treatment effect into a simplified and interpretable way. Potential topics including adjusted or unadjusted analysis; stratified vs unstratified hazard ratio; collapsibility and subgroup; p-values; etc. We will give clinically relevant opinions and recommendations based on our interpretation and illustrate the idea using some case studies.



E9(R1) STATISTICAL PRINCIPLES FOR CLINICAL TRIALS: ADDENDUM: ESTIMANDS AND SENSITIVITY ANALYSIS IN CLINICAL TRIALS

Guidance for Industry

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> U.S. Department of Health and Human Services Food and Drug Administration Center for Drug Evaluation and Research (CDER) Center for Biologics Drahaston and Research (CDER)

> > May 2021 ICH

Revision 1

Adjusting for Covariates in Randomized Clinical Trials for Drugs and Biological Products Guidance for Industry

DRAFT GUIDANCE

U.S. Department of Health and Hannan Services Food and Drug Administration Center for Drug Evaluation and Research (CDER) Center for Biologics Evaluation and Research (CBER)

> May 2021 Biostatistics

> > Revision 1

- ICH E9 encourages the identification of "covariates and factors expected to have an important influence on the primary variables"
- Adjusting for baseline covariates in statistical analysis of a randomized clinical trial can result in more efficient use of the data
 - Compensate for the chance imbalance
 - Improve efficiency in treatment effect estimation
 - Focus on prognostic baseline factors
- ICH E9(R1) requests a precise description of the treatment effect reflecting the clinical questions posed by the trial objective
 - Estimation for a treatment effect should align with the estimand

Vocabulary







Three Myths

MYTH # I: Marginal and conditional estimands can be compared directly

MYTH # 2: Standardization methods using adjusted model can guarantee improvement in precision by addressing marginal estimand

MYTH # 3: Conditional estimand by stratified analysis for time-to-event data may NOT be improved by adjusting additional prognostic covariates

Myth # I: Marginal and Conditional Estimands Can BeiGene be Compared Directly

Linear model

- Marginal estimand coincides with conditional estimand
- $\bullet \quad Y = \beta_0 + \beta_1 Z + \beta_2 X$
- Efficiency gain by reducing residual variance if the covariates are prognostic

Non-linear model

- Marginal estimand and conditional estimand differs for common efficacy measures, such as
 - logit(Pr(Y = 1)) = $\beta_0 + \beta_1 Z + \beta_2 X$ for binary endpoint;
 - $\lambda(t) = \lambda_0(t) \exp(\beta_1 Z + \beta_2 X)$ for time-to-event endpoint
- Adjusted estimator associated with a larger variance and further away from the null

Marginal Odds Ratio Differs from Conditional Odds BeiGene Ratio Due to Non-Collapsibility

	Percentage of	Succe	ss rate	Odds ratio	
	target population	New drug	Placebo		
Males	50%	80.0%	33.3%	8.0	
Females	50%	25.0%	4.0%	8.0	
Combined	100%	52.5%	18.7%	4.8	

- Treatment effect in each subgroup defined by gender are identical, OR=8 (conditional)
- Treatment effect in the combined population is different, OR=4.8 (marginal)

FDA Guideline. (2021), Adjusting for Covariates in Randomized Clinical Trials for Drugs and Biological Products Guidance for Industry

Marginal Odds Ratio Numerically Moves Away from the Conditional Odds Ratio As the Prognostic Effect Deepens

- Two subgroups $(s_1 \text{ and } s_2)$ with equal prevalence
- Odds ratio is constant within each subgroup (OR=3)
- In control arm
 - $Pr(Y = 1|s_1) = 0.1; Pr(Y = 1|s_2)$ varies in [0.1, 0.9]
- In treated arm
 - $Pr(Y = 1|s_1)$ and $Pr(Y = 1|s_2)$ are derived through the constant OR
- Marginal odds ratio is calculated through

 $\frac{\Pr(Y = 1 | \text{treated}) / \Pr(Y = 0 | \text{treated})}{\Pr(Y = 1 | \text{control}) / \Pr(Y = 0 | \text{control})}$





A Real Example with Feedback from FDA

- Primary estimand uses the marginal odds ratio $\frac{p_1}{1-p_1}/\frac{p_0}{1-p_2}$
- Primary analysis uses the logistic regression with baseline covariates
 - Regression coefficient as the estimate of the primary estimand
- FDA comments
 - Estimand uses the marginal odds ratio, but the logistic regression uses the conditional odds ratio, which does not align with the estimand
 - Clarify whether the population-level summaries are marginal odds ratios or conditional odds ratios, and provide adequate clinical justifications for these choices
 - For conditional odds ratios, the definitions of the estimands should include the variables (and their transformations) on which the odds ratios will condition
 The choice of estimand should inform the choice of analysis approach, not the reverse





MYTH # 2: Standardization Methods using Adjuster ^{BeiGene} Model Can Guarantee Improvement in Precision Addressing Marginal Estimand

Standardization (estimating standardized outcome distributions using covariate specific estimates of the outcome distribution)

- I. Model fitting: fit a regression model (e.g., GLM) considering treatment and pre-specified baseline covariates
- 2. Prediction: predict potential outcomes under treatment and under control for each patient
- 3. Average: average treatment effect on predicted outcome
- 4. Obtain a proper standard error (e.g.. Sandwich estimator, delta method or bootstrap)

BeiGene Simulation Studies to Understand Potential Benefit of Standardization

Study I: Logistic regression – odds ratio

Study II: Cox regression – hazard ratio

Simulation



Data generation

- Generalized linear model on binary outcome: $Pr(Y = 1 | Z = z, X = x) = g^{-1}(\beta_0 + \beta_1 z + \beta_2 x)$
 - Z is treatment; $X \sim N(0,1)$ is baseline covariate
 - $g^{-1}(\cdot)$ is the link function
- Number of simulations = 1000
- **Target treatment effect**: marginal treatment effect
- Analysis method
 - Logistic regression without covariate adjustment
 - Logistic regression with covariate adjustment + standardization
 - The analysis models are misspecified
- Performance measure
 - Robustness of standardization: bias, SE
 - Efficiency of standardization: relative SE



Robustness of Standardization

- Data generation P(Y = 1 | Z = z, X = x) = exp(-|-1.8 + 1.6z 0.6x|)
- Summary measure: odds ratio (OR)

	N=500	N=1000	N=2000
Method	Bias(SE)	Bias(SE)	Bias(SE)
Unadjusted estimator	0.32(1.84)	0.20(1.22)	0.06(0.83)
Standardization	0.33(1.78)	0.21(1.20)	0.06(0.81)

N is total sample size

 The standardized estimator is consistent when generalized linear model (with canonical link functions) is mis-specified in randomized trials



Efficiency of Standardization

Data generation

- $P(Y = 1 | Z = z, X = x) = \Phi(-1.8 + 1.6z + \beta_2 x),$
- $\beta_2 = 0, 0.5, 1, 1.5, 2, 4, 6, 8$
- ► N=1000

β_2	0	0.5	1	1.5	2	4	6	8
Relative SE (standardization/unadjusted estimator)	1.00	0.95	0.85	0.79	0.70	0.50	0.43	0.39

 Adjusting for prognostic covariates improves efficiency for estimating unconditional treatment effect



Challenges on Time-to-event Outcome

- Proportional hazard assumption can only hold for either the analysis model for unconditional treatment effect or the analysis model for conditional treatment effect, and censoring distribution adds further complexity
- Application of standardization method to estimate a marginal causal hazard ratio is not straightforward, an analytical solution might be intractable (Daniel et al. 2020)
- Semiparametric estimators (Lu and Tsiatis 2008) are complex and its properties need further exploration
- One could consider alternative summary measures such as restricted mean survival time (RMST) or survival probability difference, however such experience in clinical trials is limited

Simulation Study by Daniel et al (2020)



Data generation

• $\lambda(t) = \gamma \lambda t^{\gamma-1} \exp(\theta \cdot z + \beta_x x);$ • Uniform enrollment over 2 years • Administrative censoring at 10

years

Scenarios

- $\theta = 1, \beta_x = 0$: treatment effect, null covariate effect
- $\theta = 1, \beta_x = 1$: treatment effect, covariate effect

Treatment effect, Covariate Effect (Coefficients:(1,1)) Treatment effect, Null covariate Effect (Coefficients:(1,0)) 6-8 -6-Density 4 -2-12 09 1.1 1.2 08 0,7 0.8 0.4 06 1.0 10 log(HR) log(HR) Methods ! Unadjusted ! Adjusted marginal ! Conditional

- The simulation approach to obtain marginal treatment effect is computationally expensive. Need to carefully consider study time frame and censoring distribution
- Efficiency gain is not guaranteed for a given study

 $\lambda = 0.1, \gamma = 1.5$



Marginal Estimand Can be Helpful

- To enhance comparability apple to apple comparisons across data sources
 - especially, when leveraging non-randomized data
- Estimating a marginal effect by standardization approach can in general lead to efficiency gain
 - The efficiency gain can be sizable when the covariate effect is strong
 - Enhance understanding of the strength of covariate effect before applying the approach with pre-specification
- Interpretation and numerical results is based on the pre-defined population of interest





Myth # 3: Conditional Estimand can NOT be Improved by Adjusting Additional Prognostic Covariates

- Covariate adaptive randomization (e.g., stratified permuted block randomization) is commonly used
- For study with time-to-event endpoints, the primary analysis is often a stratified analysis following the stratified randomization
- Factors used in randomization is usually a subset of potential prognostic covariates
- Is there any room to improve for efficiency without losing robustness?





Original Article

Robust tests for treatment effect in survival analysis under covariate-adaptive randomization

Ting Ye, Jun Shao 🔀

First published: 19 August 2020 | https://doi.org/10.1111/rssb.12392 | Citations: 5

A Simulation to Evaluate the Efficiency of Covariates BeiGene Adjusted Analysis in a Randomized Study

- Scenario I: stratification factor in randomization is subset of prognostic factors
 - Data generation model: $\lambda(t) = \lambda_0 \exp(\theta \cdot z + \beta_x x_1 + \beta_x x_2^2)$
 - $\lambda_0 = \log(2) / 12$; *z*: treatment indicator
 - Negative θ favors treatment
 - Prognostic factors: $x_1 = 0$ or 1 with $Pr(x_1 = 1) = 0.5$; $x_2 \sim U(0,1)$
 - Randomization: stratified by x_1
- Scenario 2: continuous prognostic factor is discretized for randomization
 - Data generation model: same as scenario 1
 - Randomization: stratified by x_1 and discretized x_2 ($x_2 \le 0.5$ and $x_2 > 0.5$)
- Scenario 3: event time is not from the Cox type model
 - Data generation model: $T = \exp(\theta \cdot z + \beta_x x_1 + \beta_x x_2^2) + \epsilon$
 - z, x_1, x_2 are same as scenario I; $\epsilon \sim EXP(1)$
 - Positive θ favors treatment
 - Randomization: stratified by x_1 and discretized x_2 ($x_2 \le 0.5$ and $x_2 > 0.5$)



Testing Methods

Log-rank tests

- Unstratified log-rank test
 - Type I error is conservative under stratified randomization
- Stratified log-rank test
 - Type I error is robust under stratified randomization
 - Only account for discrete covariates

Tests based on Cox model:

$$\lambda(t) = \lambda_0(t) \exp(\alpha \cdot z + \beta_1 x_1 + \beta_2 x_2)$$

- Robust score test (Ye and Shao 2020)
 - Type I error is robust to stratified randomization and model misspecification
 - Inefficient if the analysis model is very different from the true model
- Wald test
 - Type I error is inflated if analysis model is wrong

Scenario I: Stratification Factor in Randomization is BeiGene Subset of Prognostic Factors



Testing methods

- Unstratified Log-rank test
- Log-rank test stratified by x₁
 - Robust score test adjusting both x_1 and continuous x_2
- Wald test

Observations

- The unstratified Log-rank test is conservative under stratified randomization
- Power is enhanced by adjusting the covariate not considered in randomization (i.e., stratified log rank test and robust score test)
- Wald test performs well because the working model is close to true model

Scenario 2: Continuous Prognostic Factor is Discretized for Randomization



- Robust score test adjusting x1 and continuous x2 - Wald test

Testing methods

- Unstratified Log-rank test
- Log-rank test stratified by x₁ and discrete x₂
- Robust score test adjusting both x_1 and continuous x_2
- Wald test

►

►

Observations

Adjusting for more prognostic factors in stratified log-rank test can enhance power

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Scenario 3: Event Time is not Simulated from a Cox Type Model



Testing methods

- Unstratified Log-rank test
- Log-rank test stratified by x_1 and discrete x_2
- Robust score test adjusting x_1 and continuous x_2
- Wald test

Observations

- Stratified log-rank test is superior to the robust score
- Type I error of Wald test is severely inflated since working model is very wrong



Robust score test adjusting x1 and continuous x2 --- Wald test ----



What Works Well for the Time to Event Analysis

- Unstratified log-rank test tests a marginal treatment effect. It is conservative under stratified randomization
- For additional prognostic factors not part of stratification factors for randomization, including these variables into the analysis model may further enhance the study power
- Some covariates are continuous in nature, adjusting these covariates using their continuous scale may help to improve efficiency
- "All models are wrong", consider a robust approach to draw valid statistical inference
- Stratified log-rank test performs well in general







Conclusions

- Estimation of treatment effect should align with the target estimand
- Conditional estimand and marginal estimand are both population level summary of treatment effect and should be clearly differentiated
- For binary outcome, standardization procedure is a robust approach to estimate the marginal treatment effect
 - Efficiency gain can be expected if the adjusted covariates have strong prognostic effect
- For time-to-event outcome, standardization procedure is tricky to implement, and complicated by censoring and time frame
- Stratified log-rank test performs well, robust score test approaches offer good promises



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